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Abstract

In this study, it is tried to give a system of Markov chains approach to the trend of annual mean temperatures in Turkey. For this reason, a data of annual mean temperatures between the years 1965 - 2012 of 58 meteorological stations of Turkey are used. Each scenario is given as a solution to a quadratic programming problem for which is spanned by the transition matrices of twelve groups. The steady-state distribution method given here facilitates the multiple station Markov chain applications. In the meantime, the linear regression approaches in which the averages of station groups are considered as independent variables are as well introduced. It is also given comments on some scenarios while an extreme scenario as a solution to a few problems is pointed out to which is feared in respect of climatic change.

Keywords: Annual mean temperatures, climatic change, Markov chain, Aggregation by steady-state distribution, Linear regression

Introduction and Motivation

The climate is the slowly varying aspects of atmosphere-hydrosphere-land the surface system, according to the American Meteorological Society. Climate is used for a statistical description, for example, such as the average temperature or precipitation over a period of time which measures the variations on these quantities (Dessler, 2012, p.1). Briefly, the climate is the average weather conditions that do not change for many years in a very large area.

Climate change is any systematic change in the long-term statistics of climate elements such as temperature, precipitation, pressure or wind speeds over several decades or longer, again by the American Meteorological Society (Dessler, 2012, p.4).

An important part of climate change studies is about temperature changes. According to our search on Scholar.Google (April 24, 2017), about one seventh of all climate change studies involve precipitation changes. The same research points out in about 4% of climate change studies the temperature changes and precipitation changes take place together.

There are a number of studies made on the annual mean temperature changes in Turkey. Dogan et al. (2015) on the trend directions by backward- and forward-shifted methods, Kadioglu (1997) on the trend analysis by the Mann-Kendall test, Ustaoglu (2012) on the trend analysis by the Mann-Kendall test for the North-west of Turkey, Can and Atimtay (2002) on the trend and abrupt changes by the Mann-Kendall and Wald-Wolfowitz tests, Karaburun et al. (2012) on the trend of annual mean temperatures for Marmara region by the Mann-Kendall test and Theissen polygons, Tayanc et al. (1998) on the inhomogeneity of annual mean temperatures in Turkey by the Wald-Wolfowitz and Kruskal-Wallis tests, among others. As it is known, the Mann-Kendall and Wald-Wolfowitz tests are nonparametric tests. Among these techniques, there are some other methods, known as Theissen polygons, the inverse distance

weights and kriging. The first two methods are based on the catchment geometry (Ball and Luk, 1998), and the third method is known as a method of interpolation for spatial data. These methods are generally used for the rainfall and temperature data. For the details and history of kriging see Stein (1999).

In this study, we used the steady-state distribution of Markov chains to aggregate the stations. Since kriging has a stochastic spatially dependent component, in a sense, our approach resembles kriging. In essence, the purpose of this study is to give some explanations of considered models and catch a direction to climate change by comparing the station groups in terms of their frequencies included in the models.

In average temperature changes, the increases and decreases are primarily striking. As the explained reasons in the next section, let us denote the increases in mean temperatures in successive years by 1 and the decreases or unchanges by 0. The transition matrix estimated from the mean of annual mean temperature changes of 58 meteorological stations for 48 years, which is the subject of our research, by the Maximum Likelihood method is

$$P = \begin{bmatrix} 0.28571 & 0.71429\\ 0.60000 & 0.40000 \end{bmatrix},$$

and its steady-state distribution is

$$\pi = \begin{bmatrix} 0.4565 & 0.5435 \end{bmatrix}.$$

Looking only at these results, it can be worried about climate change, because according to the steady-state distribution, it is expected that in the long run 45.65% of all times the average temperatures will decrease or not change, and increase 54.35%. In other words, in the average temperatures it would be expected a decrease in every 1.84 years, even though it is expected to decrease or remain unchanged at about every 2.19 years, in the long term. And this transition matrix will again be considered in Section 4. By these comments, we can be pessimistic. In a sense, we will discuss some details that can be called as scenarios in the following sections.

Methodology and Models

Since the climate system is one of the most complex systems that scientists have tried to come to grips with, there are many approaches to understanding this system and new approaches are still ongoing. One of these approaches is to take advantage of the simulated system by generating as much system dynamics as possible in powerful computers. But the most powerful supercomputers have not also been able to capture all the processes of system dynamics. So, instead of complex numerical models, scientists are directed to simple conceptual models. Stochastic models are such a conceptual tool to understand better the climate system (Imkeller and von Storch, p.vi). Since for the climate system there are, in general, lots of data obtained from many stations, Markov chain and point process models are similar to the extent that they are restricted to single-station applications, and are not easily generalizable to multiple station applications. References to the Markov chains and related derived models can be seen in (von Storch and Navarra, 1995, p.201). In the case of an irreducible or ergodic Markov chain, it is appropriate to consider the steady-state distribution from the asymptotic analysis point of view (Clark, 2008, p.46). Taking this in mind, it is tried to give a new model to deal with multiple stations through steady-state distribution.

Aggregation and decomposition to reduce the state space are two general techniques to solve large scale Markov decision processes (Boucherie and van Dijk, 2017, p.372). Aggregation has different kinds of purposes, such as state aggregation and time aggregation. In other words, the key idea in aggregation is to lead to an effective reduction of dimensionality (Yin and Zhang, 2013, p.53). However, in contrast to state aggregation, which generally results in an approximate model due to the loss of Markov property, time aggregation suffers no loss of accuracy, because the Markov property is preserved. (Cao, et al., 2002). For the theory,

applications and details of aggregation in Markov models see (Anisimov, 2008, Chapters 8-10), and the references therein.

For this study, the data of annual mean temperatures between the years 1965 - 2012 of 58 meteorological stations of Turkey are used. As there are few data such as 48 years, it is observed both the autocorrelation coefficients of the each station and of each of the station groups up to the 12th lag are smaller than the AR (1) (β_1) coefficient in respect of their absolute values. For this reason, and the emphasis that temperature changes can be modeled by Markov chains in many cases in Dobrovolski (2000), we decided to use a homogeneous Markov chain model consisting

of a state space 0 and 1. In this Markov chain, the cases are taken as the state 0 if the difference of mean temperatures between consecutive years is negative or zero, and as the state 1, if the difference is positive. Thus, it can be said that the state 0 represents the case the mean temperatures do not change or decrease and the state 1 represents the case the mean temperatures increase. Under this state determination, the transition matrices for each of the 58 stations are estimated by the Least Squares method and then their steady-state distributions are found. By the observation on the steady-state distributions it was determined that 58 stations have 12 different steady-state distributions up to 4th decimal place.

Climatic Zone		
(Geographical	Stations Name	Station ID
Region)		
Black Sea	Amasya, Artvin, Bartin, Bayburt,	17085, 17045, 17020, 17089,
	Bolu, Duzce, Giresun, Gumushane,	17070, 17072, 17034, 17088,
	Kastamonu, Ordu, Rize Samsun,	17074, 17033, 17040, 17030,
	Sinop, Tokat	17026, 17086
Marmara	Bilecik, Bursa Edirne, Kirklareli,	17120, 17116, 17050, 17052,
	Kocaeli, Yalova	17066, 17119
Central Anatolia	Aksaray, Ankara, Cankiri, Karaman,	17192, 17130, 17080, 17246,
	Kayseri, Kirsehir, Nevsehir, Sivas,	17196, 17160, 17193, 17090,
	Yozgat	17140
Aegean	Afyonkarahisar, Aydin, Denizli,	17190, 17234, 17237, 17220,
	Izmir, Kutahya, Mugla, Usak	17155, 17292, 17188
Mediterranean	Adana, Antakya, Burdur, Isparta,	17351, 17372, 17238, 17240,
	Kahramanmaras, Mersin	17255, 17340
Eastern Anatolia	Ardahan, Bingol, Elazig, Erzincan,	17046, 17203, 17201,17094,
	Igdir, Malatya, Mus, Tunceli, Van	17100, 17199, 17204, 17165,
		17172
South East	Adiyaman, Gaziantep, Kilis,	17265, 17261, 17262, 17275,
Anatolia	Mardin, Sanliurfa, Siirt	17270, 17210

Table 1. Meteorological stations distribution by climatic zones (geographical regions).

Station Group No	Included Station(s) Name	Steady-state Distribution
1	Adiyaman, Afyonkarahisar, Bartin, Bayburt, Bolu, Gaziantep, Igdir, Isparta, Kirklareli, Malatya, Mus, Ordu, Sivas	[0.5652 0.4348]
2	Aksaray, Antakya, Ardahan, Bilecik, Bingol, Erzincan, Kayseri, Kilis, Rize, Siirt, Usak, Yozgat	[0.5870 0.4130]
3	Adana, Amasya, Burdur, Cankiri, Düzce, Elazig, Gumushane, Karaman, Kutahya, Mardin, Yalova	[0.5435 0.4565]
4	Ankara, Artvin, Bursa, Edirne, Kastamonu, Mugla, Samsun, Sinop	[0.5217 0.4783]
5	Aydin, Denizli, Giresun, Kocaeli	[0.5000 0.5000]
6	Nigde, Sanliurfa, Van	[0.6304 0.3696]
7	Kirsehir, Nevsehir	[0.6087 0.3913]
8	Tokat	[0.5869 0.4131]
9	Tunceli	[0.5594 0.4406]
10	Kahramanmaras	[0.5250 0.4750]
11	Mersin	[0.6153 0.3847]
12	Izmir	[0.4783 0.5217]

Table 2. Station groups by their stations and steady-state distributions.

The estimated transition matrices of 12 station groups are as follows:

$$P_{1} = \begin{bmatrix} 0.41420 & 0.58580 \\ 0.76154 & 0.23846 \end{bmatrix},$$

$$P_{2} = \begin{bmatrix} 0.42593 & 0.57407 \\ 0.81579 & 0.18421 \end{bmatrix},$$

$$P_{3} = \begin{bmatrix} 0.37455 & 0.62545 \\ 0.74459 & 0.25541 \end{bmatrix},$$

$$P_{4} = \begin{bmatrix} 0.328125 & 0.671875 \\ 0.732950 & 0.267050 \end{bmatrix},$$

$$P_{5} = \begin{bmatrix} 0.33696 & 0.66304 \\ 0.66304 & 0.33696 \end{bmatrix},$$

$$P_{6} = \begin{bmatrix} 0.49425 & 0.50575 \\ 0.86275 & 0.13725 \end{bmatrix},$$

$P_7 = \begin{bmatrix} 0.48214 \\ 0.80556 \end{bmatrix}$	0.51786 0.19444]
$P_8 = \begin{bmatrix} 0.29630\\ 0.94737 \end{bmatrix}$	0.70370 0.05263
$P_9 = \begin{bmatrix} 0.40000\\ 0.76190 \end{bmatrix}$	0.60000 0.23810],
$P_{10} = \begin{bmatrix} 0.44000\\ 0.61905 \end{bmatrix}$	0.56000 0.38095],
$P_{11} = \begin{bmatrix} 0.44828\\ 0.88235 \end{bmatrix}$	0.55172 0.11765]
$P_{12} = \begin{bmatrix} 0.27273\\ 0.61905 \end{bmatrix}$	0.72727 0.38095]

	Station No											
Lag	1	2	3	4	5	6	7	8	9	10	11	12
1	50727	49839	51344	54797	51726	52228	57953	55488	42870	33529	36948	48254
2	0.06676	0.01455	0.02005	0.08271	0.06822	0.03880	0.09453	0.06346	06210	18069	0.05137	0.05681
3	10495	07746	03376	08238	14187	08302	00489	0.00632	12521	05923	0.00799	11842
4	0.02718	0.02449	01401	0.02524	0.10781	0.03974	08603	05501	0.05401	0.06498	08638	0.01507
5	0.18561	0.22253	0.19923	0.13041	0.08457	0.21110	0.22106	0.14546	0.27684	0.24975	0.10616	0.04987
6	24215	24862	22715	09526	10326	26643	22963	14108	29952	22720	19368	0.01772
7	0.03461	0.03321	0.03026	04413	03345	0.06399	0.10467	0.01056	0.06867	12054	0.08219	11550
8	01393	02430	02059	02102	02140	00952	07902	04267	05037	0.10300	03158	01392
9	0.07693	0.10023	0.17467	0.1814	0.18505	0.03456	0.19800	0.18207	0.05888	0.07506	0.08904	0.29851
10	04491	07911	18569	23429	24001	00345	21792	19250	0.06998	01565	00076	35263
11	0.06995	0.06444	0.11300	0.23135	0.24472	0.00850	0.13946	0.15369	05958	01173	03272	0.32198
12	01230	0.01228	0.01361	17646	16039	.018980	00528	07056	0.13039	0.08090	0.05023	30476

Table 3. Autocorrelations for station groups up to lag 12.

In the next step of model determination, the stations with the same steady-state distribution are collected in the same station group and the transition frequencies observed for each station are combined to estimate the transition matrices of these station groups.

The 12 station groups respectively consist of 1st group 13 stations, 2nd group 12 stations, 3rd group 11 stations, 4th group 8 stations, 5th group 4 stations, 6th group 3 stations, 7th group 2 stations and the following groups, namely 8, 9, 10, 11 and 12th groups only 1 station; this can be observed from Table 2. The first three of these station groups consist of all of 7 different geographical regions of Turkey, the 4th station group consists of 4 different regions, the 5th and 6th station groups consist of 3 different regions and the 7th and later station groups consist of stations from only 1 geographical region. In the 1st station group there are 4 stations from the Black Sea region and in the 2nd station group there are 3 stations from the Eastern Anatolia and Central Anatolia as in the highest frequencies. And in the 3rd station group there are 3 stations from the Black Sea region and in the 4th station group there are 4 stations from the Black Sea region as in the highest frequencies.

Since the magnitude and even the direction of change in many meteorological variables is uncertain, in this study it is developed and studied many numbers of models as possible with different approaches. Each of these models which use a quadratic programming technique can be generally expressed as follows (Floudas and Visweswaran, 1995):

$$Maxf(\underline{x}) = c^T \underline{x} - \frac{1}{2} \underline{x}^T Q \underline{x}$$

subject to (ST)

$$A\underline{x} \le \underline{b}, \quad \underline{x} \ge \underline{0},$$

where \underline{x}^{1} is the transpose of vector \underline{x} , and the order relations in vectors and matrices are componentwise. As seen, a quadratic programming problem consists of a quadratic objective function and linear constraints. In our approach, the right-hand side constants of linear constraints, that is, the upper bounds are important because of the size of average temperature variability and the uncertainty of its direction of change. The models given as a quadratic programming problem can be explicitly formulated as follows:

$$Min \sum_{i=1}^{12} x_i^2 \text{ or (sometimes } Max \sum_{i=1}^{12} \sum_{j \neq i}^{12} x_i x_j \text{)}$$

ST
$$\sum_{i=1}^{12} x_i P_i \le (\ge) P_j; j = 1,...,12$$
$$\sum_{i=1}^{12} x_i = 1$$
$$x_i \ge 0; i = 1,2,...,12$$

where the variable x_i denotes the weight of i th station group and P_i denotes the Least Square Estimation of transition matrix of i th station group. As mentioned in the following, the different models have been achieved by

increasing assumptions and consequently increasing the number of constraints. Among these models, according to the assumptions used, there are the models giving precedence and not-giving precedence to the number of stations in station groups, the model bounded

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by the transition matrix closest to all transition matrices and the model bounded by the symmetric transition matrix closest to all the transition matrices in the constraint, the models derived from the first two models by adding the constraints with the squares and fourth powers of transition matrices.

Table 4 shows the number of non-zero solutions that the station groups rank among

and the order of sizes of these solutions in quadratic programming models. As an example, the station group 1 having 13 stations ranks among 45 non-zero solutions in which 36 of them are ≥ 0.05 and 10 of them are in the first order. Similarly, the station group 12 having only one station ranks among 49 non-zero solutions in which 35 of them are ≥ 0.05 and 21 of them are in the first order.

Station Group No	# of Stations	# of non-zero solutions	# of solutions ≥ 0.05	# of solutions in the first order
1	13	45	36	10
2	12	42	31	5
3	11	47	36	5
4	8	49	35	6
5	4	47	38	8
6	3	47	36	12
7	2	42	35	6
8	1	41	22	5
9	1	43	30	1
10	1	49	36	15
11	1	39	30	3
12	1	49	35	21

Table 4. Non-zero solution types related to station groups.

Station Group No	Solutions < 0.05		Solutions ≥ 0.05		Mean Weights of		
	Min [*]	Max	Min	Max	> 0	≥ 0.05	
1	0.00094	0.04344	0.05347	0.82278	0.15125	0.18076	
2	0.00094	0.04689	0.05347	0.99998	0.09509	0.11742	
3	0.00001	0.04689	0.05205	0.60236	0.07823	0.09440	
4	0.00000	0.04689	0.05128	0.78002	0.10377	0.13309	
5	0.01410	0.04689	0.05205	0.99994	0.09865	0.14947	
6	0.00001	0.04689	0.05205	1.00000	0.16274	0.20596	
7	0.00008	0.04388	0.05205	0.99999	0.11039	0.12840	
8	0.00000	0.04852	0.05084	1.00000	0.12821	0.21986	
9	0.00094	0.04689	0.05346	0.75281	0.07996	0.10362	
10	0.00001	0.04689	0.05148	1.00000	0.12468	0.16192	
11	0.00001	0.04689	0.05151	0.69623	0.12098	0.14966	
12	0.00004	0.04863	0.05219	1.00000	0.25892	0.34932	

*The 0.00000 values on the table are actually non-zero values such that <0.00001.

Here, it is obvious that the comments to be made by considering the station numbers in the station groups will be very striking and show the significance of each station group, and hence the importance of stations in this group in the annual mean temperature changes. From Table 5, as an example, the station group 5 having 4 stations has the minimum value 0.0140 and the maximum value 0.04689 in the non-zero solutions <0.05,

while it takes at least 0.05205 and at most 0.99994 values in the non-zero ≥ 0.05 , and the corresponding solutions means of these solutions are 0.09865 and 0.14947 respectively. Similarly, station group 12 having only one the station has the minimum value 0.00004 the maximum value 0.04863 in and the non-zero solutions < 0.05, while it takes at least 0.05219 and at most1.00000 values in the non-zero solutions ≥ 0.05 and the corresponding means of these solutions are 0.25982 and 0.34932, respectively.

It is compared the weight estimates obtained by two different sampling methods over the actual observed values for the changes in annual mean temperatures for errors. The first sampling was performed on the ranked data and the second on the time ordered data, i.e., over the combined historical data. The comparisons are made over the Mean Absolute Deviation Error (MADE) and Mean Square Error (MSE). For illustrating the diversity of models, only the best three and worst last of them are given in detail in Tables 6 and 7.

Table 6. The MSE values and number of non-zero weights < 0.05 and ≥ 0.05 for the models of ranked and ordered data.

Model [*]	Ranked Data MSE	# of non-zero weights	Model	Ordered Data MSE	# of non- zero weights
MINSF3 - U	0.511840	8(4)**	MINSR9 - U	0.187396	3(2)
MINCO12 - U	0.512015	9(9)	GM - A	0.187396	12(12)
MINCO3 - U	0.512454	9(6)	MAXCPSYM - A	0.190957	7(5)
MINCAB***	0.187396	12(12)	MINS5 - MINSF5	0.191590	4(1)
			MIN5 - U	0.191590	7(5)
			MINCO5 - A	0.191590	7(7)
			MIN8 - A	0.194454	1
			MINS8 - AU	0.194454	1

* Here A and U point out auto scaled and unscaled techniques for QSB program, respectively.

**The frequencies in the parentheses belong to the solutions ≥ 0.05 .

***Here, in essence GM - A and MINCAB models give the same solution in quadratic programming problem.

Table 7. The MADE values and number of non-zero weights < 0.05 and ≥ 0.05	for the models of
ranked and ordered data.	

Model	Ranked Data MSE	# of non- zero weights	Model	Ordered Data MSE	# of non-zero weights
MINSF4 - U	2.475107	4(2)	MINSR9 - U	1.002128	3(2)
MINCO12 - U	2.480284	9(9)	GM - A	1.002128	12(12)
MINSF3 - U	2.480780	8(4)	MINCO12 - U	1.019858	9(9)
MINCAB	1.002128	12(12)	MAXCPSYM - A	1.020284	7(5)
			MIN8 - A	1.075461	1
			MINS8 - AU	1.075461	1

These models are given in detail as follows:

MINSF3:	MINCO12:	MINCO3:	MINSR9:
$Min\sum_{i=1}^{12}x_i^2$	$Min\sum_{i=1}^{12}x_i^2$	$Min\sum_{i=1}^{12}x_i^2$	$Min\sum_{i=1}^{12}x_i^2$
ST	ST	ST	ST
$\sum_{i=1}^{12} x_i P_i^2 \le P_3^2$	$\sum_{i=1}^{12} x_i P_i \le P_{12}$	$\sum_{i=1}^{12} x_i P_i \le P_3$	$\sum_{i=1}^{12} x_i P_i \ge P_9$
$\sum_{i=1}^{12} x_i P_i^4 \le P_3^4$	$x_i \ge x_{i+1}; i = 1,,8$ $x_7 \ge x_j; j = 10,11,12$	$x_i \ge x_{i+1}; i = 1,,8$ $x_7 \ge x_j; j = 10,11,12$	$\sum_{i=1}^{12} x_i P_i^2 \ge P_9^2$
$\sum_{i=1}^{12} x_i = 1.$	$\sum_{i=1}^{12} x_i = 1.$	$\sum_{i=1}^{12} x_i = 1.$	$\sum_{i=1}^{12} x_i = 1.$
GM:	MAXCPSYM	M: MINS5:	MINSF5:
$Min\sum_{i=1}^{12}x_i^2$	$Max\sum_{i=1}^{12}\sum_{j\neq i}^{12}$	$x_i x_j$ $Min \sum_{i=1}^{12} x_i^2$	$Min\sum_{i=1}^{12}x_i^2$
ST	ST	ST	ST
$\sum_{i=1}^{12} x_i P_i \le \overline{P}, \overline{P} = \sum_{i=1}^{12} x_i = 1.$	$\sum_{i=1}^{12} P_i / 12 \qquad \sum_{i=1}^{12} x_i P_i$ $\sum_{i=1}^{12} x_i =$	$\leq P_{SYM} \qquad \sum_{i=1}^{12} x_i P_i \leq P_S \sum_{i=1}^{12} x_i P_i^2 \leq P_S \sum_{i=1}^{12} x_i P_S \sum$	$\sum_{i=1}^{12} x_i P_i \le P_5$ $\sum_{i=1}^{5} \sum_{i=1}^{12} x_i P_i^2 \le P_5^2$ $\sum_{i=1}^{12} x_i P_i^4 \le P_5^4$ $\sum_{i=1}^{12} x_i = 1.$

MIN5:

MINCO5:

MINSF4:

MINCAB:

 $Min\sum_{i=1}^{12} x_i^2 \qquad Min\sum_{i=1}^{12} x_i^2$

 $Min\sum_{i=1}^{12} x_i^2$

 $Min\sum_{i=1}^{12} x_i^2$

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ST

ST

$$\begin{split} \sum_{i=1}^{12} x_i P_i &\leq P_5 & \sum_{i=1}^{12} x_i P_i \leq P_{12} & \sum_{i=1}^{12} x_i P_i^2 \leq P_4^2 \\ \sum_{i=1}^{12} x_i P_i &\leq x_{i+1}; i=1,...,7 & \sum_{i=1}^{12} x_i P_i^4 \leq P_4^4 & \sum_{i=1}^{12} x_i P_i \leq P_{AB} \\ \sum_{i=1}^{12} x_i &= 1. & \sum_{i=1}^{12} x_i = 1. & \sum_{i=1}^{12} x_i = 1. & \sum_{i=1}^{12} x_i = 1. \end{split}$$

MIN8:

ST

MINS8:

$$Min \sum_{i=1}^{12} x_i^2 \qquad Min \sum_{i=1}^{12} x_i^2$$

ST ST ST
$$\sum_{i=1}^{12} x_i P_i \le P_8 \qquad \sum_{i=1}^{12} x_i P_i^2 \le P_8^2$$

$$\sum_{i=1}^{12} x_i = 1. \qquad \sum_{i=1}^{12} x_i = 1.$$

ST

where the matrix P_{SYM} in the model MAXCPSYM is a symmetric transition matrix which is the closest to all the transition matrices of station groups. In other words, the matrix P_{SYM} can be easily obtained by the following model MINAB with b = a, i.e., as the matrix P_{AA} :

MINAB:

$$Min\sum_{i=1}^{12} (P_i - P_{AB})^2, P_{AB} = \begin{bmatrix} a & 1-a \\ 1-b & b \end{bmatrix}$$

ST

$$a \le 1, b \le 1, a \ge 0, b \ge 0.$$

The models MINSF3, MINCO12 and MINCO3 can be interpreted as minimizing the square of weights over the squares and fourth powers of transition matrices as constraints

with P_3 and P_{12} , and additionally giving the precedence to the number of stations in station groups in the last two models, i.e., in the models MINCO12 and MINCO3. In the solutions with extra powers of transition matrices, it has been observed some slight corrections relative to models without extra powers. In the model MINSF3, it is obvious that there is no precedence to the number of stations in station groups. Similarly, the other models can be interpreted. Here, it should not be forgotten that in the comments to be made for each model, it would be very useful to consider station numbers in station groups; because a station group in the solution including stations from different climatic zones will give it extra meaning and importance. It is also useful to give such a detail that the quadratic programming problem does not provide a feasible solution when the MINAB model has

an extra constraint $a \leq b$ on a and b values.

In this case, the quadratic programming problem gives positive solutions which do not satisfy

$$\sum_{i=1}^{12} x_i = 1$$

constraint. By dividing the coefficients from zero by the sum

$$\sum_{i=1}^{12} x_i$$

that is, with normalization relatively usable solutions can be obtained.

Linear Regression Approaches

In this section, it is investigated the variation of the annual mean temperatures in Turkey by the linear regression approach. For this reason, it should be based on the assumption that they have approximately normal distribution whereas the increases and decreases observed in annual mean temperatures have nonnormal distributions as shown in the graphs below.

As seen from Table 8, the adjusted \mathbf{p}^2

 R_{adj}^2 determination coefficients, R_{adj}^2 's show that the regression models on the single station group means are also valid such that the single station group means explain at least 57.5, and at most 98.4, percent of the variation of annual mean temperatures. In fact, it is easy to see that there is a positive correlation between the number of stations in the station groups and the coefficients of determination for the regression models on single station group means, which is exactly 0.5740.

	R_{adj}^2		R_{adj}^2
$\hat{Y}_i = -0.107793 + 0.980618G_{1,i}$	0 9698	$\hat{Y}_i = 0.006012 + 0.885888G_{2,i}$	0 9767
(0.023025) (0.025788)	0.7070	(0.019989) (0.020419)	0.9707
$\hat{Y}_i = 0.009964 + 0.947467G_{3,i}$	0.9838	$\hat{Y}_i = 0.009881 + 1.02966G_{4,i}$	0 8972
(0.016670) (0.018150)	0.7050	(0.041947) (0.051967)	0.0772
$\hat{Y}_i = -0.007066 + 1.12893G_{5,i}$	0.0727	$\hat{Y}_i = -0.000063 + 0.740424G_{6,i}$	0.0126
(0.046589) (0.064085)	0.8/3/	(0.038675) (0.034147)	0.9120
$\hat{Y}_i = 0.019806 + 0.640916G_{7,i}$	0.8662	$\hat{Y}_i = 0.022189 + 0.730136G_{8,i}$	0.0205
(0.047825) (0.037545)	0.8002	(0.036470) (0.031614)	0.9203
$\hat{Y}_i = 0.022590 + 0.638294G_{9,i}$	0 7012	$\hat{Y}_i = 0.045251 + 0.853630G_{10,i}$	0 7301
(0.071469) (0.062106)	0.7012	(0.066783) (0.075607)	0.7571
$\hat{Y}_i = 0.036170 + 0.889269G_{11,i}$	0 5749	$\hat{Y}_i = 0.001616 + 1.08271G_{12,i}$	0.6915
(0.085237) (0.113990)	0.3749	(0.0726998) (0.107815)	0.0713

Table 6. Regression estimates of annual mean temperature changes by the station group mean

Distribution	Kolmog Smirn	OrOV OV	Anders Darlir	son 1g	Chi-Squared			
	Statistic	Rank	Statistic	Rank	Statistic	Rank		
Burr (4P)	0.03231	2	1.9646	3	71.011	4		
Dagum (4P)	0.03108	1	1.9441	2	99.126	5		
Gen. Logistic	0.03281	3	1.8495	1	55.527	3		
Normal	0.06518 15		12.081	15	202.71	18		

Table 9. Goodness of fit summary for the fitted distributions.

Furthermore, assuming the correctness of the results obtained in quadratic programmingbased models, multiple linear regression models parallel to these results are also considered and the determination coefficients of these models are also found to be larger. The histograms of the increases and decreases of annual mean temperatures obtained from all stations and the distributions compiled are given in Figures 1 - 3, and the goodness of fit summary of these three distributions and the normal distribution by Easyfit program is given in Table 8 above.



Fig 1. Four-parameter Burr distribution for annual mean temperature changes. ($k = 1.7487 = \alpha,98199.0 = \beta,64411.0 = \gamma,64410.0$)



Fig 2. Four-parameter Dagum distribution for annual mean temperature changes. ($k = 0.57823 = \alpha,90.427 = \beta,40.648 = \gamma,40.204$)



Fig 3. Three-parameter generalized logistic distribution for annual mean temperature changes. ($k = 0.07453 = \sigma_{3}0.55838 = \mu_{2}0.05912$)

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Fig 4. Normal distribution for annual mean temperature changes. ($\sigma = 1.0295, \mu = -0.00977$)

In multiple linear regressions, it should be noted that C represents the constant, i.e. the intercept of the model. From the respect of coefficients in the models, a pleasing point here is that the 12th station is usually placed in the last orders in the case of multiple regressions

whereas it is usually located in the first orders with its weights in quadratic programming solutions. Here it should not be forgotten that the station (group) 12, i.e., Izmir has a tendency to have the highest average annual temperatures among all of the station groups.

			MLR	Induced [*]	R_{adi}^2
Model			Estimate	MLR Estimate	iiij
		Weight	Coefficient	Coefficient	
	С	-	02327		
	G_1	.04994	.25723	.17175	
	G_2	.03929	.17865	.26452	
MINSF3-U	G_3	.08062	.30371	.39672	.9956
	G_4	.09671	.12212	.13950	(.9952)
	G_6	.02910	.02683		
	G_7	.02028	.02543		
	G_9	.06663	.01742		
	G_{12}	.61743	.05146		

*Induced estimates consist of models which are regressed over statistically non-zero coefficients.

Model			MLR Estimate	Induced [*] MLR Estimate	R_{adj}^2	
		Weight	Coefficient	Coefficient		
	С	-	.01374			
	G_3	.07409	.58463	.71678		
MANGDO	G_4	.18581	.21492	.28625	0005	
YM- A	G_5	.22226	.13929		.9883 (.9880)	
	G_8	.08445	.07038			
	G_9	.00409	.01378			
	G_{10}	.03015	.10044	.10021		
	G_{12}	.39915	16099	14238		
		Weight	Coefficient	Coefficient		
MIN8 - A	С	-	.02219		.9205	
MINS8 - AU	G_8	1.00000	.73014	.73046	(.9217)	
		Weight	Coefficient	Coefficient		
	С	-	02197			
MINCO12- U	G_1	.08409	.24704	.17175	.9956 (.9952)	
	G_2	.08409	.25385	.26452		
	G_3	.08409	.35053	.39672		
	G_4	.08409	.10424	.13950		
	G_5	.08409	03176			
	G_{12}	.57956	.06538			
		Weight	Coefficient	Coefficient		
	С	-	00105			
MINSR9 -	G_6	.19958	.46026	.46022	.9712 (.9719)	
U	G ₉ .7528		.16611	.16615	(
	G_{12}	.04761	.34966	.34960		

Table 10-2. Multiple linear regression estimates of annual mean temperatures changes of Turkey by the station group means.

*Induced estimates consist of models which are regressed over statistically non-zero coefficients.

Model			MLR Estimate	Induced [*] MLR Estimate	R_{adj}^2
		Weight	Coefficient	Coefficient	
	С	-	02648	02012	
	G_1	.08403	.28841	.23596	.9956
MINCO3 -	G_2	.08403	.25214	.21767	(.9955)
0	G_3	.08403	.33962	.38868	
	G_4	.08403	.15487	.13655	
	G_5	.08403	07554		
	G_6	.00008	01343		
	G_7	.00008	.03930		
	G_8	.00008	05797		
	G_{12}	.57962	.06780		
		Weight	Coefficient	Coefficient	
	С		014920		
	G_1	.08333	.20243	.14224	
	G_2	.08333	.23024	.28641	
	G_3	.08333	.25262	.33662	
	G_4	.08333	.19702	.16907	0000
MINCAB GM - A	G_5	.08333	03581		.9966 (.9966)
	G_6	.08333	.02529		
	G_7	.08333	.03598		
	G_8	.08333	01808		
	G_9	.08333	.00283		
	G_{10}	.08333	.02615		
	G_{11}	.08333	.05117	.06693	
	G_{12}	.08333	.03347		

 Table 10-3. Multiple linear regression estimates of annual mean temperatures changes of Turkey by the station group means.

*Induced estimates consist of models which are regressed over statistically non-zero coefficients.

Possible Approaches over Observed a and b Values

In this section, taking into consideration the 120 possible values of transition probabilities $0 \rightarrow 0$, a and $1 \rightarrow 1$, b in the estimates of transition matrices of station groups other than those mentioned in previous sections we give some summary results. In addition, we give the models with the right hand side constraint restriction by P_3 , and by the transition probability matrix of the mean of annual mean temperature changes of all stations. The reason why we place the model with P_3 here is that the determination coefficient of the 3rd station group is larger than that of all other individual station groups. For the models based on the probabilities aand b, it is worth noting that some of these models have not feasible solutions because of the constraints of quadratic programming problem (see Fig. 4). Fig. 4 also shows the transitions between the solutions on the boundaries. For example, on the first 'row' of $G12 \leftarrow G10$ Figure 4, the notation corresponds to the transition from Solution S1 to Solution S9, where the solutions evolve from G10 to G12, although G5 is also a part of the solution sometime. It should be noted here that the indices of station groups are in the magnitude order of weights in the solutions. Here any model with no feasible solution indicates that the tendency of increase/decrease in mean temperatures will not be in the direction of this model. On the other hand, Solution S92 specifically corresponds to the MIN8-A solution, and Solution S8 roughly coincides with the GMTPM solution:

$$Min\sum_{i=1}^{12}x_i^2$$

ST

$$\sum_{i=1}^{12} x_i P_i \leq P_{GM} ,, \sum_{i=1}^{12} x_i = 1$$

where P_{GM} is the transition matrix estimated from the mean of annual mean temperature changes of all meteorological stations.

When the solutions are examined under similar thoughts, it is observed that the orders of magnitudes of the weights of station groups

change as the a values decrease from big to small values. The weights of station groups can be summarized as follows: the upper left corner is G12 corner, the upper right corner is G10 corner, the lower left corner is G8 corner and the lower right corner is G6 corner (we mean S72 solution). In a sense, similar words can be said for the edges of the feasible region: the station G12 on the upper corner of left vertical edge having the greatest weight leaves its place to the station G8 on the lower side, while the solutions are accompanied by the station groups G4 and G5 in the middle orders and the station group G5 is left later. Again in the magnitude order, on the last line of Figure 4 (the row above the solution to S92), the solution containing the station groups G6, G11, G8 and G7 transforms into a solution consisting of the station groups G8, G12 and G4 on the left corner. In the solutions on the right side of the feasible region and even to the right of the column starting with S1 (except S1), and including S40 and S41, the station groups G10, G7, G6 and G1 take place and in the solution S50 the order changes as G7, G6 and G10, and in the solutions S51 and S52 the order changes as G6, G7 and G11. In the continuation of these rows towards the low, the solution S61 includes the station groups G6, G7, G10 and G11, and the solutions S62 and S63 include G6, G7 and G11 and G6, G11 and G7 as the first three components, respectively. On the bottom row, the solutions S73, S74 and S75 contain the station groups G6, G7 and G11.

Finally, consider the solutions S35, S36, S45 and S46 at the intersection with the largest number of station groups. Using the short notations given on the sides of Figure 4, these solutions are respectively G12,5,4,10,3,9,8,1,2,7,11,6;G12,8,4,5,3,9,1,2,10,11;G12,4,5,3,9,1,2,11,10,7,6 and G8,12,4,5,3,9,1,2,11,10 Note that the solution groups S36 and S46 comprise 10 of station groups, while the solutions S35 and S45 include all station groups.

Conclusion and Some Remarks

In this study, using the observed average annual temperature values for 48 years in 58 meteorological stations of Turkey we examined some scenarios which could react on the future annual mean temperature changes. For this purpose, firstly, aggregating the Markov chains related to the annual mean temperatures of 58 stations by their steadystate distributions, 12 station groups are obtained and different quadratic programming models are considered. In these models, the right side transition matrices in the constraints indicate the direction of the temperature changes in a sense and the weights are the projections on 12 station groups which show the importance of each station group.

Similar to the application in annual mean temperatures, for the differences in mean rainfall between successive years, if we take 0 for the negative or zero cases, and 1 for the positive cases, it is also obvious that the same approach can be applied to the changes in mean precipitation. Even the Markov chain approach given here can be used to examine for the increases and decreases in mean temperatures and mean precipitations together, with a state space

 $S = \{(i, j) : i, j = 0, 1\}$

However, in such a case it may be necessary to take into account the dependence structure between mean temperatures and mean precipitations.

Some of the findings of this study are: the station groups of 4, 10 and 12 have the largest number of non-zero solutions. The station groups of 3, 5 and 6 are in second order in respect of number of non-zero solutions. The third, fourth and fifth orders pertain to the station groups 1, 9 and the station groups 2 and 7, respectively. Here, it should be noted that station groups do not take place lonely in the solutions. In these non-zero solutions, the station groups 12, 10, 6, 1, and 5 are the most

frequently seen as the first order weights (see Table 4). In all non-zero solutions, in respect of the mean weights, the first five orders pertain to the station groups 12, 6, 1, 8 and 10. In the solutions ≥ 0.05 the orders of mean weights pertain to the station groups 12, 8, 6, 1 and 10 (see Table 5).

In the quadratic programming models, a variety of solutions have been obtained, ranging from solutions containing only 1 station group to solutions containing 12 station groups with different approaches. Some of these solutions are given in Section 2. In Section 3, regression models which try to regulate the annual mean temperatures of Turkey over the station groups with non-zero weights in the quadratic programming given in Section 2 are considered. Statistically, only the coefficients of these first four station groups were significant in the regression models against the quadratic programming solutions containing the first four station groups and/or some other station groups. The adjusted determination coefficient of these regression models is as high as 0.9956. It should be noted that the first four station groups contain a total of 44 stations (75.9% of all stations).

And finally, in Section 4, some possible approaches over observed a and b values which are also related to the quadratic programming models are given in detail.

In the study, the predictions of mean temperature changes for the future are not included. That is the reason why these estimates increase the volume of the article and our changeability on how to use the weights obtained by quadratic programming. When the weights are taken as probabilistic or deterministic, it is possible to carry out different simulation studies. It may even be interesting to compare over afteasts which one of the probabilistic or deterministic approaches is closer to reality. S_i :ith feasible N:No Solution FeasibleSolution

	G12	G12,10	G12,10	G12,10	G10,12	G10,12	G12,10,5	G12,10,5	G10		1			
	\downarrow	\downarrow	\downarrow	\downarrow	\downarrow	\downarrow	\downarrow	\downarrow	\downarrow	G10,7,6,1,5		1		
	G8,12,4,5	G8- autosc	G8,4,11,2,12,3,9	G8,11,4,2,3,12,9,6	G8,11,4,2,3,12,9,6	G8,11,6,2,7,9,1,3	G8,11,6,7,2	G8,11,6,7,2	G6,11,8,7	\downarrow	G7,6,10			
										G6,11,8,7	\downarrow			
	b										G6,11,7			
												G6		
0,38095	S9	S8	S7	S6	S5	S4	S 3	S2	S1	Ν	N	N		G12 ←G10
0,33696	S19	S18	S17	S16	S15	S14	S13	S12	S11	S10	N	N		G12,8←G10,7,6,1,5
0,26705	S29	S28	\$27	S26	S25	S24	S23	S22	S21	S20	N	N		G12,8←G10,7,6,1,9,2,5,3,11
0,25541	S39	S38	\$37	S36	S35	S34	S33	S32	S31	S30	N	N		G12,8,4,5←G10,7,6,1,9,2,5,11,3
0,23810	S49	S48	S47	S46	S45	S44	S43	S42	S41	S40	N	N		G12,8,4,5←G10,7,6,1,2,11,9,3,5
0,19444	S60	S59	S58	S57	S56	S55	S54	S53	S52	S51	S50	N		G8,12,4,5←G7,6,10
0,18421	S71	S70	S69	S68	S67	S66	S65	S64	S63	S62	S61	N		G8,12,4,5←G6,7,10,11
0,13725	N	S82	S81	S80	S79	S78	S77	S76	S75	S74	S73	S72		G8,12,4←G6
0,11765	N	S91	S90	S89	S88	S87	S86	S85	S84	S83	N	N		G8,12,4←G6,11,8,7
0,05263	N	S92	N	N	N	N	N	N	N	N	N	N	а	G8-autosc
	0,272730	0,296300	0,328125	0,336960	0,374550	0,400000	0,414200	0,425930	0,440000	0,448280	0,482140	0,494250		

Fig 5. Solutions on the feasible region for the models with the observed a and b values.

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