## ECONOMIC LIBERALIZATION AND THE CAUSAL RELATIONS AMONG MONEY, INCOME, AND PRICES: The case of Pakistan

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This study re-examines the causal relations between money and the two variables, i.e., income and prices. Using annual data from 1959-60 to 2003-04, examining the stochastic properties of the variables used in the analysis, and taking care of the shifts in the series due to the start of the economic liberalization program in the early 1990s, we investigate the causal relations between real money and real income, nominal money and nominal income, and nominal money and prices. The analysis indicates, in general, the long run relationship between the money, income, and prices. The analysis further suggests a one way causation from income to money in the long run implying that real factors rather than money supply has played a major role in increasing Pakistan's national income. The study fails to find the active role of money in changing income even after taking care of possible shifts in these variables due to economic reforms. As regards the causal relationship between money and prices, the analysis suggests a uni-directional causality from money to prices, implying that monetary expansion increases inflation in Pakistan.

### I. Introduction

Money, Income, and Prices are important macroeconomic variables playing crucial role in any economy. There has been a long debate regarding their role in economics, particularly, the role of money in determination of income and prices. According to Monetarists, money plays an active role and leads to changes in income and prices. In other words, changes in income and prices in an economy are mainly caused by changes in the money stocks, i.e., the direction of causation runs from money to income and prices without any feedback. Contrary, Keynesians argues that money does not play an active role in changing income and prices. In fact, income plays the leading role in changing money stocks via demand for money, implying that the direction of causation runs from income to money without any feedback. Similarly, changes in prices are mainly caused by structural factors.

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Empirically, the two opposite views led the economists to test the causal relation of money with income and prices.

In this context, Sims (1972) developed a test of causality based on Granger approach and applied it to the U.S. data to examine the causal relation between money and income. He found the evidence of a uni-directional causality from money to income, as claimed by the Monetarists.

However, the subsequent studies on the issue did not support Sims' findings. In fact, Williams, Goodhart, and Gowland (1976) applying Sims procedure in the U.K. found a uni-directional causality from income to money, which is opposite to Sims' findings. They also found the evidence of a uni-directional causality from money to prices. Similarly, Barth and Bennett (1974), replicating Sims test in the Canadian economy, Lee and Li (1983), investigating causality among money, income, and prices in Singapore, Joshi and Joshi (1985), examining causality between money and income in India, etc., found the evidence of a bi-directional causality between income and money. Lee and Li (1983), also found a uni-directional causality from money to prices. On the other hand Brillembourg and Khan (1979), using a longer data set supported Sims' findings and found a uni-directional causality from money to income and prices in the U.S. However, Dyreyes, Starleaf, and Wang (1980), examining the pattern of causality between money and income for six industrialized countries, found a bi-directional causality in the U.S., contrary to Sims (1972), and Brillembourg and Khan (1979). Similarly, they found a uni-directional causality from money to income in Canada, contrary to Barth and Bannett (1974).

#### The Pakistani Evidence

The above discussion clearly indicate that empirical evidence regarding causal relations of money with income and prices remain inconclusive. The situation is not different in the case of Pakistan. For example, Khan and Siddiqui (1990) found unidirectional causality from income to money and bi-directional between money and prices. On the other hand, Bengali, Khan, and Sadaqat (1999) found a bi-directional causality between money and income, and uni-directional from money to prices. Abbas (1991) performing the causality test in the Asian countries found a bi-directional causality between money and income in Pakistan. Jones and Khilji (1988), analysed causal relationship between money and prices in Pakistan and found evidence of a bi-directional causality with money supply leading, while Siddiqui (1990) found a bi-directional causality between the two with prices leading. However, Hussain (2006), found the evidence of weak or breakdown of relationship between money and prices indicating adoption of an alternative strategy for the conduct of monetary policy in Pakistan.

In a multivariate framework, Ahmed (2003) investigating the causal relations among money, interest rates, prices, and output in the SAARC countries found a

bi-directional causality between money and prices in Bangladesh and Pakistan. Moreover, Ahmed's results indicate the interest rate to be a good policy variable in Bangladesh and Pakistan, and money supply in India. Similarly, Abbas and Husain (2006) used the tri-variate causality analysis among money, income, and prices and found a uni-directional causality from income to money and bi-directional between money and prices with money supply leading.

This study also attempts a comprehensive investigation of the causal relation between money and the other two variables: income and prices in Pakistan. Specifically, we investigate the causal relations between real money and real income, between nominal money and nominal income, and between nominal money and prices. In this context, we use a large data set from 1959-60 to 2003-04. Further, we take care of the stochastic properties of the variables used in the analysis and early 1990s.

The rest of the paper is organized as follows. Section II discusses the data and outlines the methodology to test the stochastic properties of the variables and their interrelationship. Section III presents the descriptive statistics regarding money, income, and prices, as well as the relationship among these variables. Sections IV, and VI examine causal relations between real money and real income, nominal money and nominal income, and nominal money and prices, respectively. Finally, Section VII presents summary and conclusions of the study.

# II. Data and Methodology

In this study the annual data from 1959-60 to 2003-04, is used to investigate the causal relations of money with income and prices in Pakistan, and to take care of the economic liberalization program started in the early 1990s, the sample is further classified into two sub-samples. Various measures aimed to move towards market-based economy have had, in general, significant impacts on the economy. Hence, Sample-1, from 1959-60 to 1990-91, covers the period prior to the start of the liberalization program whereas, Sample-2, from 1991-92 to 2003-04 represents the post-liberalization period. Similarly, in the regression analysis we include a dummy from 1991-92 onwards, to take care of the possible shift in relations among variables due to the economic liberalization program.

Gross National Product (GNP) at current prices and constant prices of 1980-81 are used as nominal and real incomes. Similarly, broad measure of money (M2) and GDP deflator with base 1980-81 are used as Money and Prices, respectively. Finally, real money is obtained by deflating M2.

The principal data source is the National Accounts of Pakistan, prepared by the Federal Bureau of Statistics. The other data source include Economic Surveys by the Finance Division, and Annual Reports by Central Bank, the State Bank of Pakistan.

We start by presenting the descriptive statistics that show the basic characteristics of the variables used in the analysis. An easy and quick way to know the relation between the two variables is to see the correlation coefficient. Similarly, the lagged correlations provide some indications regarding causal relations. Then the two correlations are reported. The formal investigation starts by examining stochastic properties of variables used in the analysis. Hence, to test the stationarity of the variables the Unit Root Test is performed on the variables. In this context, the widely used Augmented Dickey Fuller (ADF) is used. We also use Phillips-Perron (PP) tests, robust to a wide variety of serial correlation and heteroskedasticity, where the truncation lag parameters are determined, following Schwert's (1987). Next, the Engle-Granger Co-integration test is applied to explore the long run relations among the variables. Finally, the causal relationships between these variables are examined through Granger causality and/or Error Correction Models (ECM). In all cases, lag lengths are decided on the basis of minimum Final Prediction Error (FPE) and Akaike Information Criteria (AIC).

### III. Money, Income, and Prices in Pakistan

The descriptive statistics of the variables used in the analysis is presented in Table 1. The table indicates that nominal money has increased over time with an average annual expansion of around 13 per cent closely followed by nominal income that has expanded by about 12.5 per cent. On the other hand, prices increased by around 7 per cent making the real money and real income to expand by around 6 per cent and 5.5 per cent, respectively. The table also shows that the money growth variables are more volatile. It may be noted that the real money growth is the most volatile variable, whereas, the real income growth is the least volatile variable.

The table also shows the descriptive statistics for the two sub-samples. Moreover, the tests for equality of means and variances between the two sub-samples are conducted. The results indicate no significant differences in the means except in the case of real income that has gone down in the second sub-sample. On the other hand, the variances in the growth in real money and prices reduced significantly in the second sub-sample.

A preliminary indication regarding relations among money, income, and prices, can be found by looking at the correlation coefficients. Table 2 reports the coefficients showing correlations between the variables expressed in growth terms. It can be seen that nominal money is not significantly correlated with either nominal income or prices in the full sample as well as in the first sub-sample. In the post liberalization period, however, nominal money has become significantly correlated with nominal income but not with prices. The high increase in correlation coefficients of nominal money with nominal income and prices in the second period can

TABLE 1

Descriptive Statistics for Growth in Money, Income, and Prices

	Full San	nple: (1960-6	1 - 2003 - 04	-)	
	Real Money	Nominal Money	Real Income	Nominal Income	Prices
Mean Std. Dev. Observations	0.0605 0.0697 44	0.1325 0.0541 44	0.0540 0.0242 44	0.1262 0.0491 44	0.0720 0.0499 44
	Pre-liberaliz	zation: (1960	-61 – 1990-	91)	
	Real Money	Nominal Money	Real Income	Nominal Income	Prices
Mean Std. Dev. Observations	0.0590 0.0781 31	0.1292 0.0576 31	0.0601 0.0231 31	0.1304 0.0540 31	0.0702 0.0555 31
	Post-liberaliz	zation: (1991	-92 – 2003-0	04)	
	Real Money	Nominal Money	Real Income	Nominal Income	Prices
Mean Std. Dev. Observations	0.0640 0.0466 13	0.1404 0.0457 13	0.0393 0.0209 13	0.1161 0.0348 13	0.0764 0.0343 13
	Equality of	of Means and	Variances		
	Real Money	Nominal Money	Real Income	Nominal Income	Prices
Mean (t-value) Variances (F)	0.2631 2.8075**	0.6852 1.5860	2.9226** 1.2231	1.0495 2.4096	0.4486 2.6262**

be seen. Similarly, high correlation between nominal income and prices, although it has gone down in the second period, can also be noted. On the other hand, real money and real income have always been significantly correlated. Here too, the coefficient has gone up in the second period.

To look at the lagged correlations between variables (in growth terms) we now proceed to see whether money, income and prices are affected by the lagged values of their own, as well as of one another. These are shown in Table 3 for up to five period lags. The table indicate that real variables are neither affected by the lagged values of their own nor by those of the other variables. These two variables seem to be correlated only at the current level, as shown in Table 2. However, this is not the case in nominal variables where significant lagged correlations exist. The most striking feature of the table is the coefficient of correlation of the third lag of

TABLE 2

Correlations Among Money, Income, and Prices (in growth terms)

		Full Sam	ple: (1960-61	-2003-04)		
	NM	NY	DF		RM	RY
NM	1.0000			RM	1.0000	
NY	0.2502	1.0000		RY	0.4500***	1.0000
DF	0.1015	0.8870***	1.0000			
		Pre-liberali	zation: (1960-	61 – 1990-91	)	
	NM	NY	DF		RM	RY
NM	1.0000			RM	1.0000	
NY	0.1938	1.0000		RY	0.4810***	1.0000
DF	0.0477	0.9225***	1.0000			
		Post-liberal	ization: (1991-	92 – 2003-04	1)	
	NM	NY	DF		RM	RY
NM	1.0000			RM	1.0000	
NY	0.6280**	1.0000		RY	0.6517***	1.0000
DF	0.3491	0.7916***	1.0000			

nominal income in nominal money. The coefficient is amazingly high implying that money is highly affected by three years back level of income. In fact, money seems to be significantly affected by the  $2^{nd}$  and  $3^{rd}$  lags of income as well as of prices. On the other hand, income and prices do not seem to be affected by the lags of money. This suggests a one-way causation from income and prices to money. Both income and prices seems to be affected by their own 1st lags as well as by the 1st lag of the other variable. The two variables are also highly correlated at the current level as

The lagged correlations in the two sub-samples are reported in Table 4. It can be seen that the pattern of lagged correlations in the first sub-sample is exactly similar to that of the full sample, i.e., the significant effects of the 2<sup>nd</sup> and 3<sup>rd</sup> lags of income and prices on money, without any feedback, indicating a one-way causation from income and prices to money. However, in the second sub-sample representing

TABLE 3 Lagged Correlations Among Money, Income, and Prices (in growth)

_	NY	NM	DF		RY	RM
		Full Samp	le: (1960-61 – 2	2003-04)		2011
NY(-1) NY(-2) NY(-3) NY(-4) NY(-5)	0.4828*** 0.1917 0.1919 0.0375 -0.0797		0.5667*** 0.2029	RY(-1) RY(-2) RY(-3) RY(-4) RY(-5)	0.0199 0.2264 0.1148 0.1324 0.1490	0.071 -0.2889 0.0411 -0.0158 -0.1105
NM(-1) NM(-2) NM(-3) NM(-4) NM(-5)	0.2139 0.1083 -0.0451 -0.0433 -0.0398	0.2489 0.0845 0.1125 -0.0684 -0.2230	0.2142 0.0937 -0.0356 0.0221 -0.0437	RM(-1) RM(-2) RM(-3) RM(-4) RM(-5)	0.1226 0.0363 -0.1455 -0.1370 -0.0892	0.2390 -0.1597 -0.1319 -0.1161 -0.2761
DF(-1) DF(-2) DF(-3) DF(-4) DF(-5)	0.3870** 0.1111 0.1398 -0.0105 -0.1791	0.0487 0.4542*** 0.4924*** 0.1324 0.0979	0.4909*** 0.1684 0.0983 0.0148 -0.2351			

TABLE 4

Lagged Correlations Among Money, Income, and Prices

	NY	NM	DF		RY	RM
		Pre-liberaliza	ation: (1960-61	<u>-1990-91)</u>		
NY(-1)	0.4747**	0.0042	0.5890***	RY(-1)	-0.2614	0.063
NY(-2)	0.1473	0.4657**	0.1157	RY(-2)	0.1313	-0.377
NY(-3)	0.1540	0.7445***	0.0041	RY(-3)	-0.1622	0.102
NY(-4)	0.0201	0.2047	-0.0370	RY(-4)	0.1458	-0.006
NY(-5)	-0.0289	0.2400	-0.1920	RY(-5)	0.0250	-0.033
NM(-1)	0.1611	0.1599	0.1397	RM(-1)	0.1401	0.243
NM(-2)	0.0275	-0.0107	-0.0451	RM(-2)	0.0019	-0.243
NM(-3)	-0.1579	0.2149	-0.2150	RM(-3)	-0.2256	-0.088
NM(-4)	-0.0702	0.0041	-0.0501	RM(-4)	-0.2031	-0.097
NM(-5)	-0.0231	-0.1636	-0.0887	RM(-5)	-0.1428	-0.257
DF(-1)	0.4041**	-0.0217	0.4789**			
DF(-2)	0.1314	0.6420***	0.1115			
DF(-3)	0.2005	0.6463***	0.0522			
DF(-4)	0.0463	0.2422	0.0031			
DF(-5)	-0.0658	0.2302	-0.2027			
		Post-liberaliz	zation: (1991-92	2-2003-04)		
NY(-1)	0.3703	0.4613	0.4303	RY(-1)	0.1113	0.319
NY(-2)	0.2047	-0.1033	0.6084	RY(-2)	-0.2817	0.172
NY(-3)	0.2785	-0.0340	0.6002	RY(-3)	0.1025	0.019
NY(-4)	-0.0286	-0.3273	0.0494	RY(-4)	0.2987	0.181
NY(-5)	-0.5378	-0.5520	-0.2630	RY(-5)	-0.3359	-0.399
NM(-1)	0.5500	0.5548	0.5650	RM(-1)	0.1387	0.208
NM(-2)	0.3653	0.3252	0.4930	RM(-2)	0.2257	0.430
NM(-3)	0.3229	-0.2207	0.6013	RM(-3)	-0.0911	-0.397
NM(-4)	0.0732	-0.3262	0.2991	RM(-4)	-0.1838	-0.198
NM(-5)	-0.0253	-0.4722	0.1657	RM(-5)	-0.0848	-0.375
DF(-1)	0.3552	0.2984	0.4749			
DF(-2)	0.1064	-0.3350	0.4518			
DF(-3)	0.0306	-0.2373	0.4006			
DF(-4)	-0.1722	-0.4934	-0.0177			
DF(-5)	-0.5062	-0.4782	-0.5574			

the post liberalization period the lag effects of income and prices on money disappear. In fact, in this period money and income<sup>2</sup> are correlated at the current level.

It can be concluded from the correlation analysis that prior to start of the economic reforms, money used to play a passive role. Both income and prices appear to take one year to adjust and then start affecting money in the second and third years. In this context, income and prices also appear to affect each other in the first year. The similar pattern of correlations of income and prices with money may also be due to the correlation between income and prices that has always been high. It seems that, with the start of economic reforms the feedback mechanism from money to income has started as implied by the significant correlation between the two at current level (see Table 2).

#### Casual Relations

The formal investigation of causal relations is done within the Co-integration and Error Correction Model framework. In the first step, the variables used in the analysis are tested for the unit roots by applying both the Augmented Dickey Fuller (ADF) and the Phillips Perron (PP) tests. The results are reported in Table 5 which indicate that the variables are, in general, first differenced stationary, i.e., I(1).

The investigation for causal relation between the two variables starts by estimating the co-integrating regression suggested by Engle-Granger. If co-integration is found, the Error Correction Models are estimated, other wise, the Granger causality equations are estimated. The next three sections investigate the causal relations between real money and real income, nominal money and nominal income, and nominal money and prices.

# IV. Causality between Real Money and Real Income

Looking at the causal relation between the real variables, i.e., real money and real income. The results are reported in Table 6(a). It can be seen that the ADF and PP tests in co-integrating regression are insignificant, rejecting any long run relation between real money and real income. Similarly, the F-values in the Granger equations are insignificant, rejecting any short run causal relation between the two real variables. This suggests that real money and real income are independent of each other, both in the short and long runs. However, this result has serious implications indicating that it is futile to estimate money demand function where real income is one of the important factors in determining the demand for real money. The results seems to be affected by possible shifts in variables. We now proceed to analyse those shifts due to economic reforms undertaken in the country.

<sup>2</sup> as shown in Table 2.

TABLE 5

Unit Root Tests for Money, Income, and Prices

		ADF		
	Lev	els	First Diff	ference
	W/O Trend	W. Trend	W/O Trend	W. Trend
D. 1 Manay	-0.4896	-3.3034	-4.9573**	-4.3652**
Real Money	-2.8367	-1.0063	-6.1195**	-6.6659**
Real Income	0.3143	-3.5065	-5.0124**	-4.4882**
Nominal Money	-0.3986	-1.4550	-3.6614**	-3.7112**
Nominal Income Prices	0.0893	-2.5628	-3.5485**	-3.5577**
	PP (	(W/O Trend)		
	Le	vels	First Di	fference
	(1=3)	(1=9)	(l=3)	(l=9)
- 117	-0.214	-0.103	-4.886**	-4.763**
Real Money	-3.104**	-2.930**	-6.211**	-6.745**
Real Income	0.844	1.021	-5.014**	-4.888**
Nominal Money	-0.151	-0.162	-3.612**	-3.540**
Nominal Income Prices	0.487	0.469	-3.489**	-3.309**
		PP (W. Trend	1)	
	L	evels	First D	ifference
	(1=3)	(1=9)	(l=3)	(1=9)
2 13	-2.540	-2.152	-4.823**	-4.682**
Real Money	-0.457	-0.556	-7.325**	-7.290**
Real Income	-2.600	-2.433	-5.006**	-4.852**
Nominal Money		-1.992	-3.553*	-3.457*
Nominal Income Prices	-2.779	-2.727	-3.488*	-3.295*

TABLE 6(a)

Causality Between Real Money and Real Income

		Coint	egration (Eng	le-Granger)		
	Cons	st.	Coeff.	ADF	PP(1=3)	PP(1=9)
RM on RY	-1.3446	***	1.0350***	-1.0916	-1.3868	-1.3578
Conclusion: N	o Cointegratio	n.				
Gr	anger Causa	lity		Gran	nger Causalit	у
Lag 1	DRY	DRI	M	Lag 3	DRY	DRM
DRY(-1)	-0.0324	-0.114	19	DRY(-1)	-0.1317	-0.3483
DRM(-1)	0.0588	0.27	04	DRY(-2)	0.2671	-0.7314
F-Value	0.9173	0.05	45	DRY(-3)	0.3207	0.7292
				DRM(-1)	0.0863	0.3942*
				DRM(-2)	-0.0123	-0.0885
				DRY(-3)	-0.916	-0.1173
				F-Value	1.3129	1.3281
Conclusion: N	o Short run Ca	usality.				

## Shifts in Real Money and Real Income due to Reforms

To analyse the shift in real variables due to the economic reforms that started in the early 1990s, we introduce a dummy variable in the analysis that takes the value of one from 1991-92 onwards. The results, reported in Table 6(b), show the dummy variable in the co-integrating regression to be significant, implying significant shift in the relation between real money and real income. Moreover, the ADF and PP tests are now significant at 5 per cent level of significance, indicating the existence of a long run relation between the real variables. The error term in money equation is significant at 10 per cent verifying, although weak, the long run relation. The equations indicate a uni-directional causality from real income to real money in the long run and thus provides basis for estimating the money demand function. In the short run, however, the two real variables still seem to be independent of each other. It can also be seen that real money is affected by its own first lag, a result not supported by the correlation analysis.

Table 6(b)

Causality between Real Money and Real Income (reforms)

Cointegration (Engle-Granger)							
	Const.	D	Coeff.	ADF	PP(1=3)	PP(1=9)	
RM on RY	0.1372	0.2691***	0.9110***	-2.0610**	-2.3168**	-2.1077**	

Conclusion: Existence of Cointegration.

Error Co	orrection Cau	sality	Error Correction Causality			
Lag 1	DRY	DRM	Lag 2	DRY	<u>DRM</u>	
D	-0.0254**	0.0102	D	-0.0281**	-0.0298	
e(-1)	0.0267	-0.2704*	E(-1)	0.0349	-0.2668*	
DRY(-1)	-0.3365	-0.2124	DRY(-1)	-0.3708	-0.6079	
DRM(-1)	0.1101	0.3796*	DRY(-2)	-0.0755	-1.1817	
F-Value	2.8269	0.1221	DRM(-1)	0.1106	0.4370**	
			DRM(-2)	0.0386	0.0860	
			F-Value	1.4373	1.4613	

Conclusion: Unidirectional Causality from Income to Money in the long run. No Short run Causality.

Note: \*\*\*, \*\*, \* represent significance at 1%, 5%, and 10%.

### V. Causality between Nominal Money and Nominal Income

The causal relation between nominal money and nominal income is investigated here and the first set of results reported in Table 7(a). It can be seen that the PP tests in the co-integrating regression are insignificant rejecting any long run relations between the two nominal variables. However, the ADF test is significant at 10 per cent level of significance. Hence, we can say that there is a weak evidence of any long run relation between the variables. The Error Correction equations verify the weak long run relation where the error term is significant at 10 per cent in money equation. The equations indicate a weak evidence of uni-directional causality from nominal income to nominal money in the long run with no short run causal effects. If we assume no co-integration between the nominal variables the Granger equations shows evidence of income affecting money at the 2<sup>nd</sup> lag, although the F-test is not statistically significant.

TABLE 7(a) Causality between Nominal Money and Nominal Income

			and I will					
	Со	Cointegration (Engle-Granger)						
NIM on NIV	Const.	Coeff.	ADF	PP(1=3)	PP(1=9)			
NM on NY	-1.1001***	1.0156***	-1.8588*	-1.5245	-1.4510			
Conclusion: wed	ak evidence of C :				1. 1510			

Conclusion: weak evidence of Cointegration.

	Correction Ca	asanty .	Gra	nger Causali	ty
Lag 2	DNY	<u>DNM</u>	Lag 2	DNY	DNM
e(-1) DNY(-1) DNY(-2) DNM(-1) DNM(-2) F-Value	-0.0368 0.5201** -0.0116 0.0845 0.0191 0.1815	-0.2010* -0.3109 0.1245 0.2080 -0.0165 1.0607	DNY(-1) DNY(-2) DNM(-1) DNM(-2) F-Value	0.495*** -0.06 0.115 -0.009 0.371	-0.196 0.401** 0.261 -0.052 2.346

Conclusion: Weak Evidence of Unidirectional Causality from income to money.

	Correction Ca	additty	Gra	anger Causali	ty
<u>Lag 3</u> e(-1) DNY(-1)	<u>DNY</u> 0.0655	<u>DNM</u> -0.0754	Lag 3	DNY	DNM
DNY(-2) DNY(-3) DNM(-1) DNM(-2) DNM(-3) F-Value	0.5692** -0.0685 0.2093 0.0202 0.0487 -0.0953 0.1482	-0.1587 -0.0045 0.5591** 0.0339 0.0167 -0.0253 2.5031*	DNY(-1) DNY(-2) DNY(-3) DNM(-1) DNM(-2) DNM(-3) F-Value	0.504*** -0.115 0.15 0.061 0.019 -0.111 0.288	-0.097 0.097 0.520** 0.104 0.022 -0.056 4.034**

Conclusion: Unidirectional Causality from income to money at 3 years lag.

Since the lagged correlations in Table 3 also shows the significant effects of income on money at the 2<sup>nd</sup> and 3<sup>rd</sup> lags we also do the analysis for the 3<sup>rd</sup> lag. The results show that the error term in Error Correction equations has become insignificant implying no long run relation between money and income. The equations further shows the significant effects of income on money at the 3<sup>rd</sup> lag verified by F-value. Same result is shown by Granger equations, if we ignore the error term. Hence, there is evidence of a one-way causation from nominal income to nominal money, although the existence of a long run relation between the two nominal variables is not clear. There is also a persistent evidence of nominal income affected by its own first lag, as well as affecting nominal money at the 3<sup>rd</sup> lag, and thus, verifying the results shown by the lagged correlations.

## Shifts in Nominal Money and Nominal Income due to Reforms

The results of the analysis of shifts in nominal variables due to the economic reforms are reported in Table 7(b). The co-integrating regression shows significant shift in relation between nominal money and nominal income. It also indicate the existence of a long run relation between the nominal variables as the ADF and PP tests, which are now significant at 5 per cent level of significance. However, the long run relation is not verified by the Error Correction Model where the error term is not significant in both the equations, even at the 10 per cent level of significance. The error term in money equation is significant at 11 per cent which may be considered as a weak evidence of a uni-directional causality from nominal income to nominal money in the long run. Assuming no co-integration, as in the previous case, the Granger equations show the evidence of income affecting money at the 2<sup>nd</sup> lag with F-test, not statistically significant. Similarly, the analysis for the 3<sup>rd</sup> lag provides the same result, that is, income affecting money at three years lag. Hence, there is persistent evidence of income affected by its own first lag and affecting money at the third lag, without any feed back from money.

## VI. Causality between Nominal Money and Prices

Finally, we investigate the causal relation between nominal money and prices. The first set of results is reported in Table 8(a). The ADF and PP tests are highly significant indicating the existence of a long run relation between money and prices in Pakistan. The error correction equations suggest a uni-directional causality from money to prices in the long run, and thus, support the monetarists preposition. In the short run, the two variables seem to be independent of each other. There is, however, some evidence of prices affecting money at the 2<sup>nd</sup> lag. Once again, as in the case of nominal income, we do the analysis for the 3<sup>rd</sup> lag because the lagged correlations in Table 2 indicate the significant effects of prices on money at the 2<sup>nd</sup>

Table 7(b)

Causality between Nominal Money and Nominal Income (reforms)

	Cointegration (Engle-Granger)								
	Const.	$\underline{\mathbf{D}}$	Coeff.	ADF	PP(1=3)	PP(l=9)			
NM on NY	-0.4596***	0.2896***	0.9581***	-2.1835**	-2.4117**	-2.1568**			
Conclusion: Ev	vidence of Coin	tegration.							

Error Correction Causality			Granger Causality		
Lag 2	DNY	<u>DNM</u>	Lag 2	DNY	DNM
D	-0.0158	-0.0037	D	-0.0151	0.00
e(-1)	-0.1744	-0.3066			0.00
DNY(-1)	0.5306**	-0.3211	DNY(-1)	0.4786***	-0.1872
DNY(-2)	0.1561	-0.0023	DNY(-2)	-0.0742	0.4085**
DNM(-1)	0.1022	0.2038	DNM(-1)	0.1238	0.2565
DNM(-2)	0.0330	-0.0178	DNM(-2)	0.0091	-0.062
F-Value	0.3013	0.9704	F-Value	0.4551	2.3654

Conclusion: Weak Unidirectional Causality from Income to Money in the long run. No Short run Causality.

Error Correction Causality			Granger Causality		
Lag 3	DNY	<u>DNM</u>	Lag 3	DNY	DNM
D	-0.0044	0.0165	D	-0.0144	0.0120
e(-1)	-0.3833	0.0494			0.0120
DNY(-1)	0.6903***	-0.0509	DNY(-1)	0.4839***	-0.0799
DNY(-2)	0.2349	0.1248	DNY(-2)	-0.1239	0.1045
DNY(-3)	0.4314	0.7006**	DNY(-3)	0.1388	0.5300**
DNM(-1)	-0.0252	-0.0071	DNM(-1)	0.0721	0.0948
DNM(-2)	0.0568	0.0060	DNM(-2)	0.0316	0.012
DNM(-3)	-0.0505	-0.0267	DNM(-3)	-0.1018	-0.0633
F-Value	0.0753	2.7024*	F-Value	0.2851	4.1219**

Conclusion: Unidirectional Causality from income to money at 3 years lag.

TABLE 8(a)

Causality between Nominal Money and Prices

	C	ointegration (I	Engle-Granger)		
	- Const.	Coeff.	ADF	PP(1=3)	PP(1=9)
NM on DF	3.8497***	1.6967***	-3.6957***	-2.6873***	-2.4772**
Conclusion: Ev	vidence of Cointeg	ration.			
Error C	Correction Caus	ality			

Lag 2	DDF	DNM		
e(-1)	-0.3139***	-0.0714		
DDF(-1)	0.5895***	-0.3492		
DDF(-2)	0.2160	0.4964*		
DNM(-1)	0.1626	0.1672		
DNM(-2)	0.0029	0.0451		
F-Value	0.8978	2.4457		

Conclusion: Unidirectional causality from money to prices in the long run.

Note: \*\*\*, \*\*, \* represent significance at 1%, 5%, and 10%.

and 3<sup>rd</sup> lags. The results (not reported here) show the uni-directional causality from money to prices in the long run with no short run causal effects.

### Shifts in Money and Prices due to Reforms

The results for the analysis of shifts in nominal variables due to the economic reforms are reported in Table 8(b). The dummy variable in the co-integrating regression shows no significant shift in the relation of money and prices. The results remain the same, that is, a unidirectional causality from money to prices in the long run with no causal relation in the short run.

#### VII. Summary and Conclusions

The objective of this study is to re-examine the causal relations between money and the two variables, i.e., income and prices in a comprehensive manner. Using large data set from 1959-60 to 2003-04, the stochastic properties of variables used in the analysis are examined; and noting of the shifts in the series due to reforms,

TABLE 8(b)

	Causal	ity betwe	en Nominal	Money and	Prices (rofe	
		Cointe	gration (Eng	le-Granger)	Trees (refor	ms)
NM on DF	<u>Const.</u> 3.9204***	D	Coeff.	ADF	PP(1=3)	PP(1=9)
Conclusion: F	3.9204	0.0546	1.6780***	-3.7578***	-2.7160***	-2.5096***

Conclusion: Evidence of Cointegration.

Lag 2 D e(-1). DDF(-1) DDF(-2) DNM(-1) DNM(-2) F-Value	DDF -0.0079 -0.3072*** 0.5840*** 0.2081 0.1630 0.0054 0.8839	<u>DNM</u> -0.0025 -0.0617 -0.3486 0.5148* 0.1683 0.0486 2.4032	
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Conclusion: Unidirectional causality from money to prices in the long run.

Note: \*\*\*, \*\*, \* represent significance at 1%, 5%, and 10%.

we investigate the causal relations between real money and real income, nominal money and nominal income, and between nominal money and prices.

The descriptive statistics show much lower expansion in prices, relative to money and income. Moreover, the expansions in money and income seem close to each other. The correlation analysis shows significant correlation between real money and real income but not of nominal money with either nominal income or prices. There is, however, evidence of a strong correlation between nominal variables, money and income, during the period of economic reforms. Money and prices never seem to be correlated with each other.

The lagged correlation analysis seems to suggest that prior to the start of the economic reforms, money used to play a passive role. Both income and prices appear to take one year to adjust and then start affecting money, in the second and the third years. In this context, income and prices also appear to affect each other in the first year. The similar pattern of correlations of income and prices with money may also be due to the correlation between income and prices that has always been high. It seems that with the start of economic reforms the feedback mechanism from money to income has started.

The formal analysis, however, does not verify the feedback mechanism from money. Though the economic reforms of the 1990s caused significant shifts in the relations between money and income both in real and nominal terms, money supply still seems to play a passive role in increasing national income. The analysis indicates the existence of a long run relation between money and income when expressed in real terms with income as the leading variable that affects money in the long run. On the other hand, when the two variables are expressed in nominal terms the existence of a long run relation between the two is not clear. Nevertheless, there is sufficient evidence showing income as the leading variable. In the short run, the two variables, i.e., money and income appear to be independent of each other whether expressed in real or nominal terms.

It may be mentioned here that our finding of a one-way causation from income to money is in line with previous studies conducted in Pakistan in particular, and in the developing countries, in general. Hence, it can be concluded that in these countries real factors, rather than money supply, have played a major role in increasing the National Income. On the other hand, regarding the money-prices relation, in contrast to the previous studies on Pakistan which generally find a two way causation between money and prices; our results find a one way causation from money to prices. However, the study finds the indication of a feedback mechanism from prices in the short run where they seem to affect money with two years lags. However, the previous studies are generally based on monthly analysis. Moreover, as pointed out by Hussain (2006), there appears to be a break in the money-price relationship in Pakistan. Nevertheless, it can be concluded that there exist a long run relation between money and prices with money playing the lead role.

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